# The Value of Implicit Guarantees 

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## 1. Introduction



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Why do we care?

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Crisis management: Bailouts and guarantees vs. free market economy

Crisis prevention: Regulatory approaches

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$\triangleright$ Moral hazard

- Reduced funding costs bear negative incentives to higher leverage and excessive risk-taking

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- Wealth transfer from taxpayers to creditors

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- Structural change in which "default" is no longer perceived as the same event across debt and equity markets

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Crisis prevention: Regulatory approaches
$\triangleright$ Standalone credit risk

- Better gauge of financial health than observed CDS price (cf. Hart and Zingales, 2009)

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Crisis prevention: Regulatory approaches
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- Better gauge of financial health than observed CDS price (cf. Hart and Zingales, 2009)
$\triangleright$ Taxation
- Bank levy based on funding advantage backed out of debt prices net of guarantees


## The questions

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Q. Which financial institutions benefit from these subsidies, to what extent, and how are these gains split up between shareholders and creditors?
Q. How do guarantees influence the financing strategy of banks?

## Preview of the results

$\triangleright$ Findings point to a significant funding advantage of major banks during the crisis, that is less pronounced or even inexistent for non-banks

- Structural break in the pricing assumptions for U.S. bank debt
- Stock-implied default risk estimates exceed their CDS counterparts by 1000 bps
- Effect is transitory and prices tend to converge after 2008
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- \$129bn in the case of shareholders
- \$236bn in the case of debtholders
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- \$129bn in the case of shareholders
- \$236bn in the case of debtholders
$\triangleright$ In the course of the interventions, U.S. banks shifted to fixed-rate short-term financing to exploit their TBTF status


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$\Rightarrow$ Contrast default risk as explicitly priced in the CDS market to the default risk as it is implicitly priced in the stock market (Schweikhard and Tsesmelidakis, 2012)
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$\Rightarrow$ Contrast default risk as explicitly priced in the CDS market to the default risk as it is implicitly priced in the stock market (Schweikhard and Tsesmelidakis, 2012)
$\triangleright$ Exploit the divergence between the model-implied and actual CDS prices and adjust for counterparty risk to derive the funding advantage financial institutions enjoy from being TBTF

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## Default barrier


$\triangleright$ Stochastic default barrier, which is only revealed at default

- Barrier $B=L D$, where $L \sim L N(\bar{L} ; \lambda)$
$\Rightarrow$ Increases short-term default probabilities by capturing the possibility of instantaneous default


## 3. Model estimations

## Predicted vs. observed CDS spreads

## Firm-level results



## 3. Model estimations

## Predicted vs. observed CDS spreads

## Sector aggregates





3. Model estimations Predicted vs. observed CDS spreads

## Relative deviations


3. Model estimations

## Relative deviations


(a) Acquisition of Bear Stearns by JPMorgan
(b) TARP
(c) Rescue package for Bank of America

## 3. Model estimations

## Counterparty risk adjustment



## 3. Model estimations

## Funding advantage


*All numbers are in basis points per annum.

## 4. Descriptives

## Bond data

|  |  | Offering Amounts |  |  |  |  | Maturities |  |  |  | Weighted Maturities |  |  |  | Trading Volumes |  |  |  |  |
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|  | $\underline{\text { Issues }}$ | FCB | VCB | ZCB | $\Sigma$ | $\frac{\text { OAO }}{T D}$ | FCB | VCB | ZCB | $\varnothing$ | FCB | VCB | ZCB | $\varnothing$ | $V_{T \leq 5 y}$ | $V_{5 y<}$ 任 $10 y$ | $V_{T>10 y}$ | $V_{\Sigma}$ | TT |
|  |  |  |  |  |  |  | Pre-Crisis Period |  |  |  |  |  |  |  |  |  |  |  |  |
| Sectors |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |
| Banks | 4,587 | 213.41 | 545.90 | 34.81 | 794.13 | 0.28 | 7.99 | 5.51 | 2.34 | 5.41 | 14.55 | 5.64 | 3.36 | 7.93 | 148.21 | 369.77 | 901.52 | 1,419.51 | 3,104 |
| Insurance | 1,292 | 50.02 | 63.09 | 0.12 | 113.22 | 0.56 | 8.71 | 9.44 | 14.96 | 8.87 | 12.10 | 16.32 | 9.52 | 14.45 | 94.20 | 132.82 | 220.86 | 447.88 | 2,613 |
| Real Estate | 91 | 35.10 | 2.46 | 0.00 | 37.56 | 0.70 | 11.12 | 6.84 | 0.00 | 10.70 | 12.48 | 3.83 | 0.00 | 11.91 | 5.50 | 36.17 | 104.62 | 146.29 | 151 |
| Others | 8 | 1.88 | 1.25 | 0.00 | 3.12 | 0.11 | 8.04 | 28.38 | 0.00 | 15.66 | 8.50 | 47.47 | 0.00 | 24.09 | 0.00 | 0.00 | 17.00 | 17.00 | 9 |
| Financials | 5,978 | 300.40 | 612.70 | 34.93 | 948.03 | 0.30 | 8.35 | 6.10 | 2.39 | 6.25 | 13.86 | 6.82 | 3.38 | 8.92 | 247.91 | 538.76 | 1,244.00 | 2,030.68 | 5,877 |
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| Sectors - - Cris Per |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |
| Banks | 5,513 | 517.45 | 255.01 | 45.06 | 817.53 | 0.29 | 2.50 | 8.54 | 2.17 | 2.96 | 5.77 | 4.78 | 2.58 | 5.28 | 692.97 | 312.51 | 425.76 | 1,431.24 | 2,923 |
| Insurance | 761 | 46.03 | 35.51 | 0.61 | 82.15 | 0.47 | 7.80 | 12.30 | 5.88 | 8.27 | 9.64 | 26.44 | 12.20 | 16.92 | 13.12 | 56.14 | 133.50 | 202.75 | 999 |
| Real Estate | 34 | 13.57 | 0.80 | 0.00 | 14.37 | 0.73 | 9.44 | 4.98 | 0.00 | 9.30 | 10.01 | 4.98 | 0.00 | 9.73 | 0.00 | 12.26 | 15.31 | 27.56 | 390 |
| Others | 12 | 7.34 | 1.95 | 0.00 | 9.29 | 0.15 | 4.38 | 2.50 | 0.00 | 4.21 | 4.49 | 1.86 | 0.00 | 3.94 | 14.32 | 4.50 | 5.00 | 23.82 | 28 |
| Financials | 6,320 | 584.39 | 293.27 | 45.68 | 923.34 | 0.30 | 3.66 | 9.02 | 2.18 | 3.64 | 6.16 | 7.38 | 2.71 | 6.37 | 720.40 | 385.41 | 579.57 | 1,685.37 | 4,340 |
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| Sectors - Post-Cris Period |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |
| Banks | 5,078 | 177.11 | 60.05 | 30.22 | 267.38 | 0.28 | 2.87 | 14.51 | 2.17 | 4.83 | 8.14 | 11.17 | 1.99 | 8.12 | 1,061.56 | 3,459.59 | 3,794.43 | 8,315.57 | 6,3375 |
| Insurance | 71 | 39.80 | 5.60 | 0.00 | 45.40 | 0.61 | 11.21 | 2.07 | 0.00 | 10.18 | 9.24 | 1.98 | 0.00 | 8.34 | 45.28 | 146.86 | 718.05 | 910.18 | 1104 |
| Real Estate | 47 | 21.99 | 0.00 | 0.00 | 21.99 | 0.76 | 10.33 | 0.00 | 0.00 | 10.33 | 9.81 | 0.00 | 0.00 | 9.81 | 0.00 | 276.25 | 363.08 | 639.32 | 690 |
| Others | 10 | 4.10 | 0.25 | 0.00 | 4.35 | 0.26 | 6.07 | 3.01 | 0.00 | 5.77 | 6.45 | 3.01 |  | 6.25 | 3.75 | 39.51 | 35.70 | 78.96 | 88 |
| Financials | 5,206 | 243.00 | 65.90 | 30.22 | 339.12 | 0.30 | 3.35 | 14.40 | 2.17 | 4.95 | 8.44 | 10.36 | 1.99 | 8.24 | 1,110.59 | 3,922.20 | 4,911.26 | 9,944.03 | 65,257 |

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| Real Estate | 91 | 35.10 | 2.46 | 0.00 | 37.56 | 0.70 | 11.12 | 6.84 | 0.00 | 10.70 | 12.48 | 3.83 | 0.00 | 11.91 | 5.50 | 36.17 | 104.62 | 146.29 | 151 |
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## Fix-to-floating ratio

Fixed-Rate to Floating-Rate Breakdown of New Bond Issues

floating rate percentage
il........ fixed rate percentage

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|  |  | Offering Amounts |  |  |  |  | Maturities |  |  |  | Weighted Maturities |  |  |  | Trading Volumes |  |  |  |  |
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|  | $\underline{\text { Issues }}$ | FCB | VCB | ZCB | $\Sigma$ | $\frac{\text { OAO }}{T D}$ | FCB | VCB | ZCB | $\varnothing$ | FCB | VCB | ZCB | $\varnothing$ | $V_{T \leq 5 y}$ | $V_{5 y<}$ 任 $10 y$ | $V_{T>10 y}$ | $V_{\Sigma}$ | TT |
|  |  |  |  |  |  |  | Pre-Crisis Period |  |  |  |  |  |  |  |  |  |  |  |  |
| Sectors |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |
| Banks | 4,587 | 213.41 | 545.90 | 34.81 | 794.13 | 0.28 | 7.99 | 5.51 | 2.34 | 5.41 | 14.55 | 5.64 | 3.36 | 7.93 | 148.21 | 369.77 | 901.52 | 1,419.51 | 3,104 |
| Insurance | 1,292 | 50.02 | 63.09 | 0.12 | 113.22 | 0.56 | 8.71 | 9.44 | 14.96 | 8.87 | 12.10 | 16.32 | 9.52 | 14.45 | 94.20 | 132.82 | 220.86 | 447.88 | 2,613 |
| Real Estate | 91 | 35.10 | 2.46 | 0.00 | 37.56 | 0.70 | 11.12 | 6.84 | 0.00 | 10.70 | 12.48 | 3.83 | 0.00 | 11.91 | 5.50 | 36.17 | 104.62 | 146.29 | 151 |
| Others | 8 | 1.88 | 1.25 | 0.00 | 3.12 | 0.11 | 8.04 | 28.38 | 0.00 | 15.66 | 8.50 | 47.47 | 0.00 | 24.09 | 0.00 | 0.00 | 17.00 | 17.00 | 9 |
| Financials | 5,978 | 300.40 | 612.70 | 34.93 | 948.03 | 0.30 | 8.35 | 6.10 | 2.39 | 6.25 | 13.86 | 6.82 | 3.38 | 8.92 | 247.91 | 538.76 | 1,244.00 | 2,030.68 | 5,877 |
|  |  |  |  |  |  |  |  |  |  | Crisi | Period |  |  |  |  |  |  |  |  |
| Sectors - - Cris Per |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |
| Banks | 5,513 | 517.45 | 255.01 | 45.06 | 817.53 | 0.29 | 2.50 | 8.54 | 2.17 | 2.96 | 5.77 | 4.78 | 2.58 | 5.28 | 692.97 | 312.51 | 425.76 | 1,431.24 | 2,923 |
| Insurance | 761 | 46.03 | 35.51 | 0.61 | 82.15 | 0.47 | 7.80 | 12.30 | 5.88 | 8.27 | 9.64 | 26.44 | 12.20 | 16.92 | 13.12 | 56.14 | 133.50 | 202.75 | 999 |
| Real Estate | 34 | 13.57 | 0.80 | 0.00 | 14.37 | 0.73 | 9.44 | 4.98 | 0.00 | 9.30 | 10.01 | 4.98 | 0.00 | 9.73 | 0.00 | 12.26 | 15.31 | 27.56 | 390 |
| Others | 12 | 7.34 | 1.95 | 0.00 | 9.29 | 0.15 | 4.38 | 2.50 | 0.00 | 4.21 | 4.49 | 1.86 | 0.00 | 3.94 | 14.32 | 4.50 | 5.00 | 23.82 | 28 |
| Financials | 6,320 | 584.39 | 293.27 | 45.68 | 923.34 | 0.30 | 3.66 | 9.02 | 2.18 | 3.64 | 6.16 | 7.38 | 2.71 | 6.37 | 720.40 | 385.41 | 579.57 | 1,685.37 | 4,340 |
|  |  |  |  |  |  |  |  |  |  | Post-Cr | isis Peri | riod |  |  |  |  |  |  |  |
| Sectors - Post-Cris Period |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |
| Banks | 5,078 | 177.11 | 60.05 | 30.22 | 267.38 | 0.28 | 2.87 | 14.51 | 2.17 | 4.83 | 8.14 | 11.17 | 1.99 | 8.12 | 1,061.56 | 3,459.59 | 3,794.43 | 8,315.57 | 6,3375 |
| Insurance | 71 | 39.80 | 5.60 | 0.00 | 45.40 | 0.61 | 11.21 | 2.07 | 0.00 | 10.18 | 9.24 | 1.98 | 0.00 | 8.34 | 45.28 | 146.86 | 718.05 | 910.18 | 1104 |
| Real Estate | 47 | 21.99 | 0.00 | 0.00 | 21.99 | 0.76 | 10.33 | 0.00 | 0.00 | 10.33 | 9.81 | 0.00 | 0.00 | 9.81 | 0.00 | 276.25 | 363.08 | 639.32 | 690 |
| Others | 10 | 4.10 | 0.25 | 0.00 | 4.35 | 0.26 | 6.07 | 3.01 | 0.00 | 5.77 | 6.45 | 3.01 |  | 6.25 | 3.75 | 39.51 | 35.70 | 78.96 | 88 |
| Financials | 5,206 | 243.00 | 65.90 | 30.22 | 339.12 | 0.30 | 3.35 | 14.40 | 2.17 | 4.95 | 8.44 | 10.36 | 1.99 | 8.24 | 1,110.59 | 3,922.20 | 4,911.26 | 9,944.03 | 65,257 |

*Monetary amounts are in billions of US\$.
**The pre-crisis period is from June 2005 to July 2007, the crisis period ranges from August 2007 to September 2009, and the post-crisis period lasts until November 2011.

## 4. Descriptives

Average Maturity of Bond Issues


## 4. Descriptives

Maturities

Maturity Structure of Bond Issues


## 4. Descriptives

## Bond data

|  |  | Offering Amounts |  |  |  |  | Maturities |  |  |  | Weighted Maturities |  |  |  | Trading Volumes |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | $\underline{\text { Issues }}$ | FCB | VCB | ZCB | $\Sigma$ | $\frac{\text { OAO }}{T D}$ | FCB | VCB | ZCB | $\varnothing$ | FCB | VCB | ZCB | $\varnothing$ | $V_{T \leq 5 y}$ | $V_{5 y<}$ 任 $10 y$ | $V_{T>10 y}$ | $V_{\Sigma}$ | TT |
|  |  |  |  |  |  |  | Pre-Crisis Period |  |  |  |  |  |  |  |  |  |  |  |  |
| Sectors |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |
| Banks | 4,587 | 213.41 | 545.90 | 34.81 | 794.13 | 0.28 | 7.99 | 5.51 | 2.34 | 5.41 | 14.55 | 5.64 | 3.36 | 7.93 | 148.21 | 369.77 | 901.52 | 1,419.51 | 3,104 |
| Insurance | 1,292 | 50.02 | 63.09 | 0.12 | 113.22 | 0.56 | 8.71 | 9.44 | 14.96 | 8.87 | 12.10 | 16.32 | 9.52 | 14.45 | 94.20 | 132.82 | 220.86 | 447.88 | 2,613 |
| Real Estate | 91 | 35.10 | 2.46 | 0.00 | 37.56 | 0.70 | 11.12 | 6.84 | 0.00 | 10.70 | 12.48 | 3.83 | 0.00 | 11.91 | 5.50 | 36.17 | 104.62 | 146.29 | 151 |
| Others | 8 | 1.88 | 1.25 | 0.00 | 3.12 | 0.11 | 8.04 | 28.38 | 0.00 | 15.66 | 8.50 | 47.47 | 0.00 | 24.09 | 0.00 | 0.00 | 17.00 | 17.00 | 9 |
| Financials | 5,978 | 300.40 | 612.70 | 34.93 | 948.03 | 0.30 | 8.35 | 6.10 | 2.39 | 6.25 | 13.86 | 6.82 | 3.38 | 8.92 | 247.91 | 538.76 | 1,244.00 | 2,030.68 | 5,877 |
|  |  |  |  |  |  |  |  |  |  | Crisi | Period |  |  |  |  |  |  |  |  |
| Sectors - - Cris Per |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |
| Banks | 5,513 | 517.45 | 255.01 | 45.06 | 817.53 | 0.29 | 2.50 | 8.54 | 2.17 | 2.96 | 5.77 | 4.78 | 2.58 | 5.28 | 692.97 | 312.51 | 425.76 | 1,431.24 | 2,923 |
| Insurance | 761 | 46.03 | 35.51 | 0.61 | 82.15 | 0.47 | 7.80 | 12.30 | 5.88 | 8.27 | 9.64 | 26.44 | 12.20 | 16.92 | 13.12 | 56.14 | 133.50 | 202.75 | 999 |
| Real Estate | 34 | 13.57 | 0.80 | 0.00 | 14.37 | 0.73 | 9.44 | 4.98 | 0.00 | 9.30 | 10.01 | 4.98 | 0.00 | 9.73 | 0.00 | 12.26 | 15.31 | 27.56 | 390 |
| Others | 12 | 7.34 | 1.95 | 0.00 | 9.29 | 0.15 | 4.38 | 2.50 | 0.00 | 4.21 | 4.49 | 1.86 | 0.00 | 3.94 | 14.32 | 4.50 | 5.00 | 23.82 | 28 |
| Financials | 6,320 | 584.39 | 293.27 | 45.68 | 923.34 | 0.30 | 3.66 | 9.02 | 2.18 | 3.64 | 6.16 | 7.38 | 2.71 | 6.37 | 720.40 | 385.41 | 579.57 | 1,685.37 | 4,340 |
|  |  |  |  |  |  |  |  |  |  | Post-Cr | isis Peri | riod |  |  |  |  |  |  |  |
| Sectors - Post-Cris Period |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |
| Banks | 5,078 | 177.11 | 60.05 | 30.22 | 267.38 | 0.28 | 2.87 | 14.51 | 2.17 | 4.83 | 8.14 | 11.17 | 1.99 | 8.12 | 1,061.56 | 3,459.59 | 3,794.43 | 8,315.57 | 6,3375 |
| Insurance | 71 | 39.80 | 5.60 | 0.00 | 45.40 | 0.61 | 11.21 | 2.07 | 0.00 | 10.18 | 9.24 | 1.98 | 0.00 | 8.34 | 45.28 | 146.86 | 718.05 | 910.18 | 1104 |
| Real Estate | 47 | 21.99 | 0.00 | 0.00 | 21.99 | 0.76 | 10.33 | 0.00 | 0.00 | 10.33 | 9.81 | 0.00 | 0.00 | 9.81 | 0.00 | 276.25 | 363.08 | 639.32 | 690 |
| Others | 10 | 4.10 | 0.25 | 0.00 | 4.35 | 0.26 | 6.07 | 3.01 | 0.00 | 5.77 | 6.45 | 3.01 |  | 6.25 | 3.75 | 39.51 | 35.70 | 78.96 | 88 |
| Financials | 5,206 | 243.00 | 65.90 | 30.22 | 339.12 | 0.30 | 3.35 | 14.40 | 2.17 | 4.95 | 8.44 | 10.36 | 1.99 | 8.24 | 1,110.59 | 3,922.20 | 4,911.26 | 9,944.03 | 65,257 |

*Monetary amounts are in billions of US\$.
**The pre-crisis period is from June 2005 to July 2007, the crisis period ranges from August 2007 to September 2009, and the post-crisis period lasts until November 2011.

## 4. Descriptives

## Bond data

|  |  | Offering Amounts |  |  |  |  | Maturities |  |  |  | Weighted Maturities |  |  |  | Trading Volumes |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | $\underline{\text { Issues }}$ | FCB | VCB | ZCB | $\Sigma$ | $\frac{\text { OAO }}{T D}$ | FCB | VCB | ZCB | $\varnothing$ | FCB | VCB | ZCB | $\varnothing$ | $V_{T \leq 5 y}$ | $V_{5 y<}$ 任 $10 y$ | $V_{T>10 y}$ | $V_{\Sigma}$ | TT |
|  |  |  |  |  |  |  | Pre-Crisis Period |  |  |  |  |  |  |  |  |  |  |  |  |
| Sectors |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |
| Banks | 4,587 | 213.41 | 545.90 | 34.81 | 794.13 | 0.28 | 7.99 | 5.51 | 2.34 | 5.41 | 14.55 | 5.64 | 3.36 | 7.93 | 148.21 | 369.77 | 901.52 | 1,419.51 | 3,104 |
| Insurance | 1,292 | 50.02 | 63.09 | 0.12 | 113.22 | 0.56 | 8.71 | 9.44 | 14.96 | 8.87 | 12.10 | 16.32 | 9.52 | 14.45 | 94.20 | 132.82 | 220.86 | 447.88 | 2,613 |
| Real Estate | 91 | 35.10 | 2.46 | 0.00 | 37.56 | 0.70 | 11.12 | 6.84 | 0.00 | 10.70 | 12.48 | 3.83 | 0.00 | 11.91 | 5.50 | 36.17 | 104.62 | 146.29 | 151 |
| Others | 8 | 1.88 | 1.25 | 0.00 | 3.12 | 0.11 | 8.04 | 28.38 | 0.00 | 15.66 | 8.50 | 47.47 | 0.00 | 24.09 | 0.00 | 0.00 | 17.00 | 17.00 | 9 |
| Financials | 5,978 | 300.40 | 612.70 | 34.93 | 948.03 | 0.30 | 8.35 | 6.10 | 2.39 | 6.25 | 13.86 | 6.82 | 3.38 | 8.92 | 247.91 | 538.76 | 1,244.00 | 2,030.68 | 5,877 |
|  |  |  |  |  |  |  |  |  |  | Crisi | Period |  |  |  |  |  |  |  |  |
| Sectors - - Cris Per |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |
| Banks | 5,513 | 517.45 | 255.01 | 45.06 | 817.53 | 0.29 | 2.50 | 8.54 | 2.17 | 2.96 | 5.77 | 4.78 | 2.58 | 5.28 | 692.97 | 312.51 | 425.76 | 1,431.24 | 2,923 |
| Insurance | 761 | 46.03 | 35.51 | 0.61 | 82.15 | 0.47 | 7.80 | 12.30 | 5.88 | 8.27 | 9.64 | 26.44 | 12.20 | 16.92 | 13.12 | 56.14 | 133.50 | 202.75 | 999 |
| Real Estate | 34 | 13.57 | 0.80 | 0.00 | 14.37 | 0.73 | 9.44 | 4.98 | 0.00 | 9.30 | 10.01 | 4.98 | 0.00 | 9.73 | 0.00 | 12.26 | 15.31 | 27.56 | 390 |
| Others | 12 | 7.34 | 1.95 | 0.00 | 9.29 | 0.15 | 4.38 | 2.50 | 0.00 | 4.21 | 4.49 | 1.86 | 0.00 | 3.94 | 14.32 | 4.50 | 5.00 | 23.82 | 28 |
| Financials | 6,320 | 584.39 | 293.27 | 45.68 | 923.34 | 0.30 | 3.66 | 9.02 | 2.18 | 3.64 | 6.16 | 7.38 | 2.71 | 6.37 | 720.40 | 385.41 | 579.57 | 1,685.37 | 4,340 |
|  |  |  |  |  |  |  |  |  |  | Post-Cr | isis Peri | riod |  |  |  |  |  |  |  |
| Sectors - Post-Cris Period |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |
| Banks | 5,078 | 177.11 | 60.05 | 30.22 | 267.38 | 0.28 | 2.87 | 14.51 | 2.17 | 4.83 | 8.14 | 11.17 | 1.99 | 8.12 | 1,061.56 | 3,459.59 | 3,794.43 | 8,315.57 | 6,3375 |
| Insurance | 71 | 39.80 | 5.60 | 0.00 | 45.40 | 0.61 | 11.21 | 2.07 | 0.00 | 10.18 | 9.24 | 1.98 | 0.00 | 8.34 | 45.28 | 146.86 | 718.05 | 910.18 | 1104 |
| Real Estate | 47 | 21.99 | 0.00 | 0.00 | 21.99 | 0.76 | 10.33 | 0.00 | 0.00 | 10.33 | 9.81 | 0.00 | 0.00 | 9.81 | 0.00 | 276.25 | 363.08 | 639.32 | 690 |
| Others | 10 | 4.10 | 0.25 | 0.00 | 4.35 | 0.26 | 6.07 | 3.01 | 0.00 | 5.77 | 6.45 | 3.01 |  | 6.25 | 3.75 | 39.51 | 35.70 | 78.96 | 88 |
| Financials | 5,206 | 243.00 | 65.90 | 30.22 | 339.12 | 0.30 | 3.35 | 14.40 | 2.17 | 4.95 | 8.44 | 10.36 | 1.99 | 8.24 | 1,110.59 | 3,922.20 | 4,911.26 | 9,944.03 | 65,257 |

*Monetary amounts are in billions of US\$.
**The pre-crisis period is from June 2005 to July 2007, the crisis period ranges from August 2007 to September 2009, and the post-crisis period lasts until November 2011.

## 4. Descriptives

## Bond data

|  |  | Offering Amounts |  |  |  |  | Maturities |  |  |  | Weighted Maturities |  |  |  | Trading Volumes |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | $\underline{\text { Issues }}$ | FCB | VCB | ZCB | $\Sigma$ | $\frac{\text { OAO }}{T D}$ | FCB | VCB | ZCB | $\varnothing$ | FCB | VCB | ZCB | $\varnothing$ | $V_{T \leq 5 y}$ | $V_{5 y<}$ 任 $10 y$ | $V_{T>10 y}$ | $V_{\Sigma}$ | TT |
|  |  |  |  |  |  |  | Pre-Crisis Period |  |  |  |  |  |  |  |  |  |  |  |  |
| Sectors |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |
| Banks | 4,587 | 213.41 | 545.90 | 34.81 | 794.13 | 0.28 | 7.99 | 5.51 | 2.34 | 5.41 | 14.55 | 5.64 | 3.36 | 7.93 | 148.21 | 369.77 | 901.52 | 1,419.51 | 3,104 |
| Insurance | 1,292 | 50.02 | 63.09 | 0.12 | 113.22 | 0.56 | 8.71 | 9.44 | 14.96 | 8.87 | 12.10 | 16.32 | 9.52 | 14.45 | 94.20 | 132.82 | 220.86 | 447.88 | 2,613 |
| Real Estate | 91 | 35.10 | 2.46 | 0.00 | 37.56 | 0.70 | 11.12 | 6.84 | 0.00 | 10.70 | 12.48 | 3.83 | 0.00 | 11.91 | 5.50 | 36.17 | 104.62 | 146.29 | 151 |
| Others | 8 | 1.88 | 1.25 | 0.00 | 3.12 | 0.11 | 8.04 | 28.38 | 0.00 | 15.66 | 8.50 | 47.47 | 0.00 | 24.09 | 0.00 | 0.00 | 17.00 | 17.00 | 9 |
| Financials | 5,978 | 300.40 | 612.70 | 34.93 | 948.03 | 0.30 | 8.35 | 6.10 | 2.39 | 6.25 | 13.86 | 6.82 | 3.38 | 8.92 | 247.91 | 538.76 | 1,244.00 | 2,030.68 | 5,877 |
|  |  |  |  |  |  |  |  |  |  | Crisi | Period |  |  |  |  |  |  |  |  |
| Sectors - - Cris Per |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |
| Banks | 5,513 | 517.45 | 255.01 | 45.06 | 817.53 | 0.29 | 2.50 | 8.54 | 2.17 | 2.96 | 5.77 | 4.78 | 2.58 | 5.28 | 692.97 | 312.51 | 425.76 | 1,431.24 | 2,923 |
| Insurance | 761 | 46.03 | 35.51 | 0.61 | 82.15 | 0.47 | 7.80 | 12.30 | 5.88 | 8.27 | 9.64 | 26.44 | 12.20 | 16.92 | 13.12 | 56.14 | 133.50 | 202.75 | 999 |
| Real Estate | 34 | 13.57 | 0.80 | 0.00 | 14.37 | 0.73 | 9.44 | 4.98 | 0.00 | 9.30 | 10.01 | 4.98 | 0.00 | 9.73 | 0.00 | 12.26 | 15.31 | 27.56 | 390 |
| Others | 12 | 7.34 | 1.95 | 0.00 | 9.29 | 0.15 | 4.38 | 2.50 | 0.00 | 4.21 | 4.49 | 1.86 | 0.00 | 3.94 | 14.32 | 4.50 | 5.00 | 23.82 | 28 |
| Financials | 6,320 | 584.39 | 293.27 | 45.68 | 923.34 | 0.30 | 3.66 | 9.02 | 2.18 | 3.64 | 6.16 | 7.38 | 2.71 | 6.37 | 720.40 | 385.41 | 579.57 | 1,685.37 | 4,340 |
|  |  |  |  |  |  |  |  |  |  | Post-Cr | isis Peri | riod |  |  |  |  |  |  |  |
| Sectors - Post-Cris Period |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |
| Banks | 5,078 | 177.11 | 60.05 | 30.22 | 267.38 | 0.28 | 2.87 | 14.51 | 2.17 | 4.83 | 8.14 | 11.17 | 1.99 | 8.12 | 1,061.56 | 3,459.59 | 3,794.43 | 8,315.57 | 6,3375 |
| Insurance | 71 | 39.80 | 5.60 | 0.00 | 45.40 | 0.61 | 11.21 | 2.07 | 0.00 | 10.18 | 9.24 | 1.98 | 0.00 | 8.34 | 45.28 | 146.86 | 718.05 | 910.18 | 1104 |
| Real Estate | 47 | 21.99 | 0.00 | 0.00 | 21.99 | 0.76 | 10.33 | 0.00 | 0.00 | 10.33 | 9.81 | 0.00 | 0.00 | 9.81 | 0.00 | 276.25 | 363.08 | 639.32 | 690 |
| Others | 10 | 4.10 | 0.25 | 0.00 | 4.35 | 0.26 | 6.07 | 3.01 | 0.00 | 5.77 | 6.45 | 3.01 |  | 6.25 | 3.75 | 39.51 | 35.70 | 78.96 | 88 |
| Financials | 5,206 | 243.00 | 65.90 | 30.22 | 339.12 | 0.30 | 3.35 | 14.40 | 2.17 | 4.95 | 8.44 | 10.36 | 1.99 | 8.24 | 1,110.59 | 3,922.20 | 4,911.26 | 9,944.03 | 65,257 |

*Monetary amounts are in billions of US\$.
**The pre-crisis period is from June 2005 to July 2007, the crisis period ranges from August 2007 to September 2009, and the post-crisis period lasts until November 2011.

## 4. Descriptives

## Trades



## 5. Capitalized subsidies

Estimate implicit subsidies resulting from the funding cost advantage.

$$
P=P V(\text { Bond })=\sum_{t=1}^{T} \frac{c N}{(1+y)^{t}}+\frac{N}{(1+y)^{T}}
$$

Shareholders' subsidies
Q. How much more would a bank have to pay (in PV terms) to raise the debt?
$\triangleright$ Re-value bond issues by increasing coupon rate to obtain non-guaranteed issue price $P_{C_{N G}}$

$$
P_{c_{G}}<P_{c_{N G}} \Rightarrow S_{c}=P_{c_{N G}}-P_{c_{G}}
$$

## 5. Capitalized subsidies

Estimate implicit subsidies resulting from the funding cost advantage.

$$
P=P V(\text { Bond })=\sum_{t=1}^{T} \frac{c N}{(1+y)^{t}}+\frac{N}{(1+y)^{T}}
$$

Shareholders' subsidies
Q. How much more would a bank have to pay (in PV terms) to raise the debt?
$\triangleright$ Re-value bond issues by increasing coupon rate to obtain non-guaranteed issue price $P_{C_{N G}}$

$$
P_{c_{G}}<P_{c_{N G}} \Rightarrow S_{c}=P_{c_{N G}}-P_{c_{G}}
$$

## Bondholders' subsidies

Q. By how much is the deterioration of bond prices offset due to the guarantee?
$\triangleright$ Re-value transactions by increasing YTM to obtain non-guaranteed transaction price $P_{y_{N G}}$

$$
P_{y_{G}}>P_{y N G} \Rightarrow S_{y}=P_{y_{G}}-P_{y N G}
$$

## 5. Capitalized subsidies

Bondholders' subsidies are estimated in two ways:

## 5. Capitalized subsidies

Bondholders' subsidies are estimated in two ways:
$\diamond$ Incremental secondary-market subsidies
$\triangleright$ Merge TRACE transaction data with issue information from Mergent FISD.
$\triangleright$ Calculated once per issue, i.e., for each reference entity that is traded between 2007-2010, select the day with the largest funding advantage and calculate the subsidy $S_{y}$.
$\triangleright$ Scale the resulting $S_{y}$ by the corresponding offering amount.
$\triangleright$ To avoid double-counting, subtract any primary-market subsidy, if there is.

## 5. Capitalized subsidies

Bondholders' subsidies are estimated in two ways:
$\diamond$ Incremental secondary-market subsidies
$\triangleright$ Merge TRACE transaction data with issue information from Mergent FISD.
$\triangleright$ Calculated once per issue, i.e., for each reference entity that is traded between 2007-2010, select the day with the largest funding advantage and calculate the subsidy $S_{y}$.
$\triangleright$ Scale the resulting $S_{y}$ by the corresponding offering amount.
$\triangleright$ To avoid double-counting, subtract any primary-market subsidy, if there is.
$\diamond$ Continuous secondary-market subsidies
$\triangleright$ Calculated daily.
$\triangleright$ Combine contemporaneous values for the funding advantage with the day-matched trading volume as inferred from TRACE.
$\triangleright$ Trading volume replaces the issue volume and gives an impression of the actual impact through time.

## 5. Capitalized subsidies

## Subsidy-to-Issue-Volume Ratios


*All numbers are in millions of US\$.

## 5. Capitalized subsidies

| Panel A - Primary Market Subsidies Implied by a Lower Coupon Rate |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: |
|  | 2007 | 2008 | 2009 | 2010 | Total |
| Banks | 3.31 | 38.25 | 77.15 | 2.58 | 121.29 |
| Insurance | 0.17 | 1.76 | 1.44 | 2.05 | 5.42 |
| Real Estate | 0.14 | 0.11 | 0.83 | 0.24 | 1.32 |
| Others | 0.00 | 0.27 | 0.86 | 0.01 | 1.14 |
| Total | 3.62 | 40.39 | 80.28 | 4.88 | 129.17 |
| Panel B - Secondary Market Subsidies Implied by a Lower Yield |  |  |  |  |  |
|  | $\underline{2007}$ | 2008 | 2009 | 2010 | Total |
| Banks | 0.47 | 93.34 | 109.13 | 0.00 | 202.94 |
| Insurance | 0.04 | 6.13 | 19.56 | 0.00 | 25.73 |
| Real Estate | 0.01 | 3.71 | 2.89 | 0.00 | 6.61 |
| Others | 0.00 | 0.27 | 0.51 | 0.00 | 0.78 |
| Total | 0.52 | 103.45 | 132.09 | 0.00 | 236.06 |
| Panel C - Overall Subsidies |  |  |  |  |  |
|  | 2007 | 2008 | 2009 | 2010 | Total |
| Banks | 3.78 | 131.59 | 186.28 | 2.58 | 324.23 |
| Insurance | 0.21 | 7.90 | 21.00 | 2.05 | 31.16 |
| Real Estate | 0.16 | 3.82 | 3.71 | 0.24 | 7.93 |
| Others | 0.00 | 0.54 | 1.36 | 0.01 | 1.91 |
| Total | 4.15 | 143.85 | 212.35 | 4.88 | 365.23 |

## 5. Capitalized subsidies

| Panel A - Primary Market Subsidies Implied by a Lower Coupon Rate |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: |
|  | 2007 | 2008 | 2009 | 2010 | Total |
| Banks | 3.31 | 38.25 | 77.15 | 2.58 | 121.29 |
| Insurance | 0.17 | 1.76 | 1.44 | 2.05 | 5.42 |
| Real Estate | 0.14 | 0.11 | 0.83 | 0.24 | 1.32 |
| Others | 0.00 | 0.27 | 0.86 | 0.01 | 1.14 |
| Total | 3.62 | 40.39 | 80.28 | 4.88 | 129.17 |
| Panel B - Secondary Market Subsidies Implied by a Lower Yield |  |  |  |  |  |
|  | $\underline{2007}$ | 2008 | 2009 | 2010 | Total |
| Banks | 0.47 | 93.34 | 109.13 | 0.00 | 202.94 |
| Insurance | 0.04 | 6.13 | 19.56 | 0.00 | 25.73 |
| Real Estate | 0.01 | 3.71 | 2.89 | 0.00 | 6.61 |
| Others | 0.00 | 0.27 | 0.51 | 0.00 | 0.78 |
| Total | 0.52 | 103.45 | 132.09 | 0.00 | 236.06 |
| Panel C - Overall Subsidies |  |  |  |  |  |
|  | 2007 | 2008 | 2009 | 2010 | Total |
| Banks | 3.78 | 131.59 | 186.28 | 2.58 | 324.23 |
| Insurance | 0.21 | 7.90 | 21.00 | 2.05 | 31.16 |
| Real Estate | 0.16 | 3.82 | 3.71 | 0.24 | 7.93 |
| Others | 0.00 | 0.54 | 1.36 | 0.01 | 1.91 |
| Total | 4.15 | 143.85 | 212.35 | 4.88 | 365.23 |

## 5. Capitalized subsidies

Secondary market

_- subsidies of financials in mn USD (left axis)
--------- subsidies of banks in mn USD (left axis)
subsidies in percent of trading volume of financials (right axis)
subsidies in percent trading volume of banks (right axis)

## 5. Capitalized subsidies

Determinants of subsidies

|  | (1) |  |  | (2) |  |  | (3) |  |  | (4) |  |  | (5) |  |  | (6) |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | Coef. | $t$ |  | Coef. | $t$ |  | Coef. | $t$ |  | Coef. | $t$ |  | Coef. | $t$ |  | Coef. | $t$ |  |
| VIX | 0.190 | 3.27 | *** | 0.196 | 3.23 | *** | 0.180 | 3.19 | *** | 0.137 | 2.41 | ** | 0.130 | 2.59 | ** | 0.126 | 2.48 | ** |
| Rating (AA) |  |  |  | 12.743 | 3.11 | *** | 12.613 | 3.09 | *** | 9.807 | 2.81 | *** |  |  |  |  |  |  |
| Rating (A) |  |  |  | 1.540 | 3.00 | *** | 1.767 | 3.81 | *** | -0.446 | -0.54 |  |  |  |  |  |  |  |
| Rating (BBB) |  |  |  | 0.458 | 1.22 |  | 1.565 | 2.75 | *** |  |  |  |  |  |  |  |  |  |
| rs | -1.343 | -0.85 |  | -1.196 | -0.72 |  | -1.763 | -1.06 |  | -6.871 | $-2.19$ | ** | -6.501 | $-2.51$ | ** | -6.100 | -2.65 | ** |
| Size | 12.116 | 5.39 | *** |  |  |  |  |  |  |  |  |  | 11.230 | 5.64 | *** | 2.824 | 2.61 | ** |
| $\beta_{r s}^{D F}$ |  |  |  |  |  |  | 4.002 | 2.47 | ** |  |  |  |  |  |  |  |  |  |
| MES |  |  |  |  |  |  |  |  |  | 144.255 | 3.37 | *** | 133.754 | 3.70 | *** | 137.554 | 3.91 | *** |
| TARP |  |  |  |  |  |  |  |  |  |  |  |  | 0.466 | 0.28 |  |  |  |  |
| TARP Amounts |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  | 0.414 | 7.80 | *** |
| Constant | -5.172 | -3.14 | *** | -5.091 | -3.12 | *** | -8.953 | -3.18 | *** | -7.442 | -3.30 | *** | -9.319 | -3.62 | *** | -8.739 | -3.44 | *** |
| Observations | 34143 |  |  | 34273 |  |  | 34273 |  |  | 23937 |  |  | 23835 |  |  | 23835 |  |  |
| Adj. $\mathrm{R}^{2}$ | 0.221 |  |  | 0.138 |  |  | 0.151 |  |  | 0.199 |  |  | 0.274 |  |  | 0.345 |  |  |
| Coef. Estimates | OLS |  |  | OLS |  |  | OLS |  |  | OLS |  |  | OLS |  |  | OLS |  |  |
| Standard Errors | CL-F |  |  | CL-F |  |  | CL-F |  |  | CL-F |  |  | CL-F |  |  | CL-F |  |  |

*Statistical significance at the $1 \%, 5 \%$, and $10 \%$ levels is denoted by ***, **, and *, respectively.

## 5. Capitalized subsidies

Determinants of subsidies

|  | (1) |  |  | (2) |  |  | (3) |  |  | (4) |  |  | (5) |  |  | (6) |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | Coef. | $t$ |  | Coef. | $t$ |  | Coef. | $t$ |  | Coef. | $t$ |  | Coef. | $t$ |  | Coef. | $t$ |  |
| VIX | 0.190 | 3.27 | + | 0.196 | 3.23 | *** | 0.180 | 3.19 | *** | 0.137 | 2.41 | ** | 0.130 | 2.59 | ** | 0.126 | 2.48 | ** |
| Rating (AA) |  |  |  | 12.743 | 3.11 | *** | 12.613 | 3.09 | *** | 9.807 | 2.81 | *** |  |  |  |  |  |  |
| Rating (A) |  |  |  | 1.540 | 3.00 | *** | 1.767 | 3.81 | *** | -0.446 | -0.54 |  |  |  |  |  |  |  |
| Rating (BBB) |  |  |  | 0.458 | 1.22 |  | 1.565 | 2.75 | *** |  |  |  |  |  |  |  |  |  |
| rs | -1.343 | -0.85 |  | -1.196 | -0.72 |  | -1.763 | -1.06 |  | -6.871 | $-2.19$ | ** | -6.501 | $-2.51$ | ** | -6.100 | -2.65 | ** |
| Size | 12.116 | 5.39 | 14.4 |  |  |  |  |  |  |  |  |  | 11.230 | 5.64 | *** | 2.824 | 2.61 | ** |
| $\beta_{r s}^{D F}$ |  |  |  |  |  |  | 4.002 | 2.47 | ** |  |  |  |  |  |  |  |  |  |
| MES |  |  |  |  |  |  |  |  |  | 144.255 | 3.37 | *** | 133.754 | 3.70 | *** | 137.554 | 3.91 | *** |
| TARP |  |  |  |  |  |  |  |  |  |  |  |  | 0.466 | 0.28 |  |  |  |  |
| TARP Amounts |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  | 0.414 | 7.80 | *** |
| Constant | -5.172 | -3.14 | *** | -5.091 | -3.12 | *** | -8.953 | -3.18 | *** | -7.442 | -3.30 | *** | -9.319 | -3.62 | *** | -8.739 | -3.44 | *** |
| Observations | 34143 |  |  | 34273 |  |  | 34273 |  |  | 23937 |  |  | 23835 |  |  | 23835 |  |  |
| Adj. $\mathrm{R}^{2}$ | 0.221 |  |  | 0.138 |  |  | 0.151 |  |  | 0.199 |  |  | 0.274 |  |  | 0.345 |  |  |
| Coef. Estimates | OLS |  |  | OLS |  |  | OLS |  |  | OLS |  |  | OLS |  |  | OLS |  |  |
| Standard Errors | CL-F |  |  | CL-F |  |  | CL-F |  |  | CL-F |  |  | CL-F |  |  | CL-F |  |  |

*Statistical significance at the $1 \%, 5 \%$, and $10 \%$ levels is denoted by ***, **, and *, respectively.

## 5. Capitalized subsidies

Determinants of subsidies

|  | (1) |  |  | (2) |  |  | (3) |  |  | (4) |  |  | (5) |  |  | (6) |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | Coef. | $t$ |  | Coef. | $t$ |  | Coef. | $t$ |  | Coef. | $t$ |  | Coef. | $t$ |  | Coef. | $t$ |  |
| VIX | 0.190 | 3.27 | *** | 0.196 | 3.23 | *** | 0.180 | 3.19 | *** | 0.137 | 2.41 | ** | 0.130 | 2.59 | ** | 0.126 | 2.48 | ** |
| Rating (AA) |  |  |  | 12.743 | 3.11 | *** | 12.613 | 3.09 | *** | 9.807 | 2.81 | *** |  |  |  |  |  |  |
| Rating (A) |  |  |  | 1.540 | 300 | *** | 1.767 | 3.81 | *** | -0.446 | -0.54 |  |  |  |  |  |  |  |
| Rating (BBB) |  |  |  | 0.458 | 122 |  | 1.565 | 2.75 | *** |  |  |  |  |  |  |  |  |  |
| rs | -1.343 | -0.85 |  | -1.196 | -0.72 |  | -1.763 | -1.06 |  | -6.871 | $-2.19$ | ** | -6.501 | $-2.51$ | ** | -6.100 | -2.65 | ** |
| Size | 12.116 | 5.39 | *** |  |  |  |  |  |  |  |  |  | 11.230 | 5.64 | *** | 2.824 | 2.61 | ** |
| $\beta_{r s}^{D F}$ |  |  |  |  |  |  | 4.002 | 2.47 | ** |  |  |  |  |  |  |  |  |  |
| MES |  |  |  |  |  |  |  |  |  | 144.255 | 3.37 | *** | 133.754 | 3.70 | *** | 137.554 | 3.91 | *** |
| TARP |  |  |  |  |  |  |  |  |  |  |  |  | 0.466 | 0.28 |  |  |  |  |
| TARP Amounts |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  | 0.414 | 7.80 | *** |
| Constant | -5.172 | -3.14 | *** | -5.091 | -3.12 | *** | -8.953 | -3.18 | *** | -7.442 | -3.30 | *** | -9.319 | -3.62 | *** | -8.739 | -3.44 | *** |
| Observations | 34143 |  |  | 34273 |  |  | 34273 |  |  | 23937 |  |  | 23835 |  |  | 23835 |  |  |
| Adj. $\mathrm{R}^{2}$ | 0.221 |  |  | 0.138 |  |  | 0.151 |  |  | 0.199 |  |  | 0.274 |  |  | 0.345 |  |  |
| Coef. Estimates | OLS |  |  | OLS |  |  | OLS |  |  | OLS |  |  | OLS |  |  | OLS |  |  |
| Standard Errors | CL-F |  |  | CL-F |  |  | CL-F |  |  | CL-F |  |  | CL-F |  |  | CL-F |  |  |

*Statistical significance at the $1 \%, 5 \%$, and $10 \%$ levels is denoted by ***, **, and *, respectively.

## 5. Capitalized subsidies

Determinants of subsidies

|  | (1) |  |  | (2) |  |  | (3) |  |  | (4) |  |  | (5) |  |  | (6) |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | Coef. | $t$ |  | Coef. | $t$ |  | Coef. | $t$ |  | Coef. | $t$ |  | Coef. | $t$ |  | Coef. | $t$ |  |
| VIX | 0.190 | 3.27 | *** | 0.196 | 3.23 | *** | 0.180 | 3.19 | *** | 0.137 | 2.41 | ** | 0.130 | 2.59 | ** | 0.126 | 2.48 | ** |
| Rating (AA) |  |  |  | 12.743 | 3.11 | *** | 12.613 | 3.09 | *** | 9.807 | 2.81 | *** |  |  |  |  |  |  |
| Rating (A) |  |  |  | 1.540 | 3.00 | *** | 1.767 | 3.81 | *** | -0.446 | -0.54 |  |  |  |  |  |  |  |
| Rating (BBB) |  |  |  | 0.458 | 1.22 |  | 1.565 | 2.75 | *** |  |  |  |  |  |  |  |  |  |
| $r_{S}$ | -1.343 | -0.85 |  | -1.196 | -0.72 |  | -1.763 | -1.06 |  | -6.871 | $-2.19$ | ** | -6.501 | $-2.51$ | ** | -6.100 | -2.65 | ** |
| Size | 12.116 | 5.39 | *** |  |  |  |  |  |  |  |  |  | 11.230 | 5.64 | *** | 2.824 | 2.61 | ** |
| $\beta_{\text {rs }}^{\text {DF }}$ |  |  |  |  |  |  | 4.002 | 2.47 | ** |  |  |  |  |  |  |  |  |  |
| MES |  |  |  |  |  |  |  |  |  | 144.255 | 3.37 | *** | 133.754 | 3.70 | *** | 137.554 | 3.91 | *** |
| TARP |  |  |  |  |  |  |  |  |  |  |  |  | 0.466 | 0.28 |  |  |  |  |
| TARP Amounts |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  | 0.414 | 7.80 | *** |
| Constant | -5.172 | -3.14 | *** | -5.091 | -3.12 | *** | -8.953 | -3.18 | *** | -7.442 | -3.30 | *** | -9.319 | -3.62 | *** | -8.739 | -3.44 | *** |
| Observations | 34143 |  |  | 34273 |  |  | 34273 |  |  | 23937 |  |  | 23835 |  |  | 23835 |  |  |
| Adj. $\mathrm{R}^{2}$ | 0.221 |  |  | 0.138 |  |  | 0.151 |  |  | 0.199 |  |  | 0.274 |  |  | 0.345 |  |  |
| Coef. Estimates | OLS |  |  | OLS |  |  | OLS |  |  | OLS |  |  | OLS |  |  | OLS |  |  |
| Standard Errors | CL-F |  |  | CL-F |  |  | CL-F |  |  | CL-F |  |  | CL-F |  |  | CL-F |  |  |

*Statistical significance at the $1 \%, 5 \%$, and $10 \%$ levels is denoted by ***, **, and *, respectively.

## 5. Capitalized subsidies

Determinants of subsidies

|  | (1) |  |  | (2) |  |  | (3) |  |  | (4) |  |  | (5) |  |  | (6) |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | Coef. | $t$ |  | Coef. | $t$ |  | Coef. | $t$ |  | Coef. | $t$ |  | Coef. | $t$ |  | Coef. | $t$ |  |
| VIX | 0.190 | 3.27 | *** | 0.196 | 3.23 | *** | 0.180 | 3.19 | *** | 0.137 | 2.41 | ** | 0.130 | 2.59 | ** | 0.126 | 2.48 | ** |
| Rating (AA) |  |  |  | 12.743 | 3.11 | *** | 12.613 | 3.09 | *** | 9.807 | 2.81 | *** |  |  |  |  |  |  |
| Rating (A) |  |  |  | 1.540 | 3.00 | *** | 1.767 | 3.81 | *** | -0.446 | -0.54 |  |  |  |  |  |  |  |
| Rating (BBB) |  |  |  | 0.458 | 1.22 |  | 1.565 | 2.75 | *** |  |  |  |  |  |  |  |  |  |
| rs | -1.343 | -0.85 |  | -1.196 | -0.72 |  | -1.763 | -1.06 |  | -6.871 | $-2.19$ | ** | -6.501 | $-2.51$ | ** | -6.100 | -2.65 | ** |
| Size | 12.116 | 5.39 | *** |  |  |  |  |  |  |  |  |  | 11.230 | 5.64 | *** | 2.824 | 2.61 | ** |
| $\beta_{r_{S}}^{\text {DF }}$ |  |  |  |  |  |  | 4.002 | 2.47 | ** |  |  |  |  |  |  |  |  |  |
| MES |  |  |  |  |  |  |  |  |  | 144.255 | 3.37 | ** | 133.754 | 3.70 | *** | 137.554 | 3.91 | *** |
| TARP |  |  |  |  |  |  |  |  |  |  |  |  | 0.466 | 0.28 |  |  |  |  |
| TARP Amounts |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  | 0.414 | 7.80 | *** |
| Constant | -5.172 | -3.14 | *** | -5.091 | -3.12 | *** | -8.953 | -3.18 | *** | -7.442 | -3.30 | *** | -9.319 | -3.62 | *** | -8.739 | -3.44 | *** |
| Observations | 34143 |  |  | 34273 |  |  | 34273 |  |  | 23937 |  |  | 23835 |  |  | 23835 |  |  |
| Adj. $\mathrm{R}^{2}$ | 0.221 |  |  | 0.138 |  |  | 0.151 |  |  | 0.199 |  |  | 0.274 |  |  | 0.345 |  |  |
| Coef. Estimates | OLS |  |  | OLS |  |  | OLS |  |  | OLS |  |  | OLS |  |  | OLS |  |  |
| Standard Errors | CL-F |  |  | CL-F |  |  | CL-F |  |  | CL-F |  |  | CL-F |  |  | CL-F |  |  |

*Statistical significance at the $1 \%, 5 \%$, and $10 \%$ levels is denoted by ***, **, and *, respectively.

## 5. Capitalized subsidies

Determinants of subsidies

|  | (1) |  |  | (2) |  |  | (3) |  |  | (4) |  |  | (5) |  |  | (6) |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | Coef. | $t$ |  | Coef. | $t$ |  | Coef. | $t$ |  | Coef. | $t$ |  | Coef. | $t$ |  | Coef. | $t$ |  |
| VIX | 0.190 | 3.27 | *** | 0.196 | 3.23 | *** | 0.180 | 3.19 | *** | 0.137 | 2.41 | ** | 0.130 | 2.59 | ** | 0.126 | 2.48 | ** |
| Rating (AA) |  |  |  | 12.743 | 3.11 | *** | 12.613 | 3.09 | *** | 9.807 | 2.81 | *** |  |  |  |  |  |  |
| Rating (A) |  |  |  | 1.540 | 3.00 | *** | 1.767 | 3.81 | *** | -0.446 | -0.54 |  |  |  |  |  |  |  |
| Rating (BBB) |  |  |  | 0.458 | 1.22 |  | 1.565 | 2.75 | *** |  |  |  |  |  |  |  |  |  |
| rs | -1.343 | -0.85 |  | -1.196 | -0.72 |  | -1.763 | -1.06 |  | -6.871 | $-2.19$ | ** | -6.501 | $-2.51$ | ** | -6.100 | -2.65 | ** |
| Size | 12.116 | 5.39 | *** |  |  |  |  |  |  |  |  |  | 11.230 | 5.64 | *** | 2.824 | 2.61 | ** |
| $\beta_{r_{S}}^{\text {DF }}$ |  |  |  |  |  |  | 4.002 | 2.47 | ** |  |  |  |  |  |  |  |  |  |
| MES |  |  |  |  |  |  |  |  |  | 144.255 | 3.37 | *** | 33.754 | 3.70 | *** | 137.554 | 3.91 | *** |
| TARP |  |  |  |  |  |  |  |  |  |  |  |  | 0.466 | 0.28 |  |  |  |  |
| TARP Amounts |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  | 0.414 | 7.80 | *** |
| Constant | -5.172 | -3.14 | *** | -5.091 | -3.12 | *** | -8.953 | -3.18 | *** | -7.442 | -3.30 | *** | -9.319 | -3.62 | *** | -8.739 | -3.44 | *** |
| Observations | 34143 |  |  | 34273 |  |  | 34273 |  |  | 23937 |  |  | 23835 |  |  | 23835 |  |  |
| Adj. $\mathrm{R}^{2}$ | 0.221 |  |  | 0.138 |  |  | 0.151 |  |  | 0.199 |  |  | 0.274 |  |  | 0.345 |  |  |
| Coef. Estimates | OLS |  |  | OLS |  |  | OLS |  |  | OLS |  |  | OLS |  |  | OLS |  |  |
| Standard Errors | CL-F |  |  | CL-F |  |  | CL-F |  |  | CL-F |  |  | CL-F |  |  | CL-F |  |  |

*Statistical significance at the $1 \%, 5 \%$, and $10 \%$ levels is denoted by ***, **, and *, respectively.

## 5. Capitalized subsidies

Determinants of subsidies

|  | (1) |  |  | (2) |  |  | (3) |  |  | (4) |  |  | (5) |  |  | (6) |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | Coef. | $t$ |  | Coef. | $t$ |  | Coef. | $t$ |  | Coef. | $t$ |  | Coef. | $t$ |  | Coef. | $t$ |  |
| VIX | 0.190 | 3.27 | *** | 0.196 | 3.23 | *** | 0.180 | 3.19 | *** | 0.137 | 2.41 | ** | 0.130 | 2.59 | ** | 0.126 | 2.48 | ** |
| Rating (AA) |  |  |  | 12.743 | 3.11 | *** | 12.613 | 3.09 | *** | 9.807 | 2.81 | *** |  |  |  |  |  |  |
| Rating (A) |  |  |  | 1.540 | 3.00 | *** | 1.767 | 3.81 | *** | -0.446 | -0.54 |  |  |  |  |  |  |  |
| Rating (BBB) |  |  |  | 0.458 | 1.22 |  | 1.565 | 2.75 | *** |  |  |  |  |  |  |  |  |  |
| rs | -1.343 | -0.85 |  | -1.196 | -0.72 |  | -1.763 | -1.06 |  | -6.871 | $-2.19$ | ** | -6.501 | $-2.51$ | ** | -6.100 | -2.65 | ** |
| Size | 12.116 | 5.39 | *** |  |  |  |  |  |  |  |  |  | 11.230 | 5.64 | *** | 2.824 | 2.61 | ** |
| $\beta_{r_{S}}^{\text {DF }}$ |  |  |  |  |  |  | 4.002 | 2.47 | ** |  |  |  |  |  |  |  |  |  |
| MES |  |  |  |  |  |  |  |  |  | 144.255 | 3.37 | *** | 133.754 | 3.70 | *** | 137.554 | 3.91 | *** |
| TARP |  |  |  |  |  |  |  |  |  |  |  |  | 0.466 | 0.28 |  |  |  |  |
| TARP Amounts |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  | 0.414 | 7.80 | *** |
| Constant | -5.172 | -3.14 | *** | -5.091 | -3.12 | *** | -8.953 | -3.18 | *** | -7.442 | -3.30 | *** | -9.319 | -3.62 | *** | -8.739 | -3.44 | *** |
| Observations | 34143 |  |  | 34273 |  |  | 34273 |  |  | 23937 |  |  | 23835 |  |  | 23835 |  |  |
| Adj. $\mathrm{R}^{2}$ | 0.221 |  |  | 0.138 |  |  | 0.151 |  |  | 0.199 |  |  | 0.274 |  |  | 0.345 |  |  |
| Coef. Estimates | OLS |  |  | OLS |  |  | OLS |  |  | OLS |  |  | OLS |  |  | OLS |  |  |
| Standard Errors | CL-F |  |  | CL-F |  |  | CL-F |  |  | CL-F |  |  | CL-F |  |  | CL-F |  |  |

*Statistical significance at the $1 \%, 5 \%$, and $10 \%$ levels is denoted by ***, **, and *, respectively.

## 6. Conclusion

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$\triangleright$ Apply a structural model framework and adjust for counterparty risk to calculate the funding advantage.
$\triangleright$ Merge with bond issue and transaction data and re-value bonds to calculate subsidies to share- and bondholders.
$\triangleright$ Capitalized subsidies amount to $\$ 365.2$ billion in total.
$\triangleright$ Banks shifted financing to short-term fixed-rate bond issues to further profit from their TBTF status.
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$\triangleright$ Banks shifted financing to short-term fixed-rate bond issues to further profit from their TBTF status.
$\triangleright$ CDS prices are biased to the downside and thus unreliable for monitoring the health of the financial system.


## A. Appendix

## Counterparty risk

## Role of counterparty risk in CDS markets

$\triangleright$ Degree to which counterparty risk affects CDS prices depends on the joint default probability of the insurer and the reference entity.
$\Rightarrow$ High in the case of contracts written on major financials as they happen to be the primary CDS dealers.
$\triangleright$ In periods of high systemic risk, both the value of guarantees (the wedge) and counterparty risk rise, moving market premiums downwards.

## A. Appendix

Role of counterparty risk in CDS markets
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$\triangleright$ In periods of high systemic risk, both the value of guarantees (the wedge) and counterparty risk rise, moving market premiums downwards.

## Adjust CDS-equity wedge for counterparty risk

1. Construct a primary dealer CDS index.
2. Measure each firm's daily beta between its CDS and the index.
3. Regress wedge on betas and control variables related to liquidity, business climate, and ratings.
4. Multiply the coefficient estimates with the beta values to obtain the counterparty risk adjustment for a given firm, maturity, and date.

## A. Appendix

## Determinants of the bond structure

|  | $T_{\text {issue }}$ |  |  |  |  |  |  |  |  | $\frac{V_{f x}}{V_{f i x}+V_{f 0}}$ |  |  |  |  |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | (1) |  |  | (2) |  |  | (3) |  |  | (4) |  |  |  | (5) |  |  | (6) |  |  |
|  | Coef. | $t$ |  | Coef. | $t$ |  | Coef. | $t$ |  | Coef |  | $t$ |  | Coef. | $t$ |  | Coef. | $t$ |  |
| $T_{\text {mat }}$ | 0.362 | 2.84 | *** | 0.298 | 1.99 | ** | 0.253 | 2.25 | ** |  |  |  |  |  |  |  |  |  |  |
| $T_{\text {issue }}$ |  |  |  |  |  |  |  |  |  |  | 0.006 | 3.41 | *** | 0.006 | 3.39 | *** | 0.01 | 2.72 | *** |
| $V_{\text {fix }}$ |  |  |  | 1.511 | 2.26 | ** |  |  |  |  |  |  |  |  |  |  |  |  |  |
| $\frac{V_{f i x}}{V_{\text {fix }}+V}$ | 3.563 | 3.25 | *** |  |  |  | 2.367 | 2.63 | *** |  |  |  |  |  |  |  |  |  |  |
| $\begin{aligned} & V_{f x}+V_{f i v} \\ & V_{\text {fixs }} \\ & \hline \end{aligned}$ |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |
| $\frac{V_{\text {fix }}}{V_{\text {mix }}+V_{\text {fict }}}$ |  |  |  |  |  |  |  |  |  |  | 0.253 | 7.85 | *** | 0.251 | 7.92 | *** | 0.16 | 5.04 | *** |
| Term Spread | -0.826 | -3.32 | *** | -0.342 | -1.07 |  |  |  |  |  | 0.059 | 6.38 | *** | 0.057 | 6.28 | *** | 0.06 | 6.73 | *** |
| Funding Adv. | -0.172 | -1.93 | * | -0.244 | -2.05 | ** |  |  |  |  | 0.027 | 4.71 | *** | 0.025 | 4.87 | *** | 0.03 | 5.28 | *** |
| Bank Dummy |  |  |  |  |  |  | -1.564 | -2.04 | ** |  |  |  |  |  |  |  | -0.11 | -4.25 | *** |
| Crisis Dummy |  |  |  |  |  |  |  |  |  |  |  |  |  | 0.206 | 5.65 |  |  |  |  |
| Post-crisis Dummy |  |  |  |  |  |  |  |  |  |  |  |  |  | 0.212 | 2.48 | ** |  |  |  |
| $V_{f i \times} \times$ Bank Dummy |  |  |  | -1.732 | -2.68 | *** |  |  |  |  |  |  |  |  |  |  |  |  |  |
| AA |  |  |  |  |  |  | 3.880 | 2.71 | *** |  |  |  |  |  |  |  |  |  |  |
| A |  |  |  |  |  |  | 4.178 | 3.03 | *** |  |  |  |  |  |  |  | 0.10 | 4.02 | *** |
| BBB |  |  |  |  |  |  | 4.212 | 2.76 | *** |  |  |  |  |  |  |  | 0.31 | 8.97 | *** |
| Constant | 5.854 | 9.11 | *** | 8.199 | 7.20 | *** | 2.330 | 1.42 |  |  | 0.296 | 10.91 | *** | 0.279 | 10.31 | *** | 0.36 | 9.77 | *** |
| Observations | 636 |  |  | 245 |  |  | 773 |  |  |  | 636 |  |  | 636 |  |  | 636 |  |  |
| Adj. $R^{2}$ | 0.062 |  |  | 0.056 |  |  | 0.046 |  |  |  | 0.197 |  |  | 0.231 |  |  | 0.29 |  |  |
| Coef. Estimates | OLS |  |  | OLS |  |  | OLS |  |  |  | OLS |  |  | OLS |  |  | OLS |  |  |
| Standard Errors | robust |  |  | robust |  |  | robust |  |  |  | robust |  |  | robust |  |  | robust |  |  |

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## A. Appendix

## Determinants of the bond structure

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| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | (1) |  |  | (2) |  |  | (3) |  |  | (4) |  |  |  | (5) |  |  | (6) |  |  |
|  | Coef. | $t$ |  | Coef. | $t$ |  | Coef. | $t$ |  | Coef. |  | $t$ |  | Coef. | $t$ |  | Coef. | $t$ |  |
| $T_{\text {mat }}$ | 0.362 | 2.84 | *** | 0.298 | 1.99 | ** | 0.253 | 2.25 | ** |  |  |  |  |  |  |  |  |  |  |
| $T_{\text {issue }}$ |  |  |  |  |  |  |  |  |  |  | 0.006 | 3.41 | *** | 0.006 | 3.39 | *** | 0.01 | 2.72 | *** |
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| $\frac{V_{f i x}}{V_{\text {fix }}+V}$ | 3.563 | 3.25 | *** |  |  |  | 2.367 | 2.63 | *** |  |  |  |  |  |  |  |  |  |  |
| $\begin{aligned} & V_{\text {fix }}+v_{\text {fig }} \\ & v_{\text {fixi }} \end{aligned}$ |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |
| $\frac{V_{\text {fix }}}{V_{\text {mix }}+V_{\text {fict }}}$ |  |  |  |  |  |  |  |  |  |  | 0.253 | 7.85 | *** | 0.251 | 7.92 | *** | 0.16 | 5.04 | *** |
| Term Spread | -0.826 | -3.32 | *** | -0.342 | -1.07 |  |  |  |  |  | 0.059 | 6.38 | *** | 0.057 | 6.28 | *** | 0.06 | 6.73 | *** |
| Funding Adv. | -0.1.72 | -1.93 | + | -0.244 | -2.05 | ** |  |  |  |  | 0.027 | 4.71 | *** | 0.025 | 4.87 | *** | 0.03 | 5.28 | *** |
| Bank Dummy |  |  |  |  |  |  | -1.564 | -2.04 | ** |  |  |  |  |  |  |  | -0.11 | -4.25 | *** |
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| A |  |  |  |  |  |  | 4.178 | 3.03 | *** |  |  |  |  |  |  |  | 0.10 | 4.02 | *** |
| BBB |  |  |  |  |  |  | 4.212 | 2.76 | *** |  |  |  |  |  |  |  | 0.31 | 8.97 | *** |
| Constant | 5.854 | 9.11 | *** | 8.199 | 7.20 | *** | 2.330 | 1.42 |  |  | 0.296 | 10.91 | *** | 0.279 | 10.31 | *** | 0.36 | 9.77 | *** |
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| Adj. $R^{2}$ | 0.062 |  |  | 0.056 |  |  | 0.046 |  |  |  | 0.197 |  |  | 0.231 |  |  | 0.29 |  |  |
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| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
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|  |  |  |  |  |  |  |  |  |  |  | 0.253 | 7.85 | *** | 0.251 | 7.92 | *** | 0.16 | 5.04 | *** |
|  | -0.826 | -3.32 | *** | -0.342 | -1.07 |  |  |  |  |  | 0.253 0.059 | 6.38 |  | 0.057 | 6.28 | *** | 0.16 | 6.73 | *** |
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| A |  |  |  |  |  |  | 4.178 | 3.03 | *** |  |  |  |  |  |  |  | 0.10 | 4.02 | *** |
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| Standard Errors | robust |  |  | robust |  |  | robust |  |  |  | robust |  |  | robust |  |  | robust |  |  |

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## A. Appendix

## Determinants of the bond structure


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## Asset dynamics

$\triangleright$ Firm assets $V$ are assumed to evolve by the diffusion

$$
\frac{d V_{t}}{V_{t}}=\mu_{V} d t+\sigma_{V} d W_{t}
$$

where $W_{t}$ is a Brownian motion, $\sigma_{V}$ the asset volatility, and $\mu v$ the drift

## B. Appendix

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$\Rightarrow$ Hence, we assume a stationary leverage, implying $\mu_{E}=\mu_{D}=\mu_{V}$
$\Rightarrow$ Pricing credit is rather about the relation between $\mu_{V}$ and $\mu_{D}$ than about $\mu_{V}$ per se, therefore set $\mu_{V}=0$ for simplicity

## B. Appendix

## Survival probability

The risk-neutral survival probability $P(t)$ that the firm value does not hit the default boundary until time $t$, i.e.,

$$
V(\tau)>L D, \quad \forall \tau<t
$$

is given by the approximate closed-form solution

$$
P(t)=\Phi\left(-\frac{A_{t}}{2}+\frac{\log (d)}{A_{t}}\right)-d \cdot \Phi\left(-\frac{A_{t}}{2}-\frac{\log (d)}{A_{t}}\right)
$$

with

$$
d=\frac{S_{0}+\bar{L} D}{\bar{L} D} \exp \lambda^{2} \quad \text { and } \quad A_{t}^{2}=\sigma_{V}^{2} t+\lambda^{2}
$$

$$
\underbrace{(1-R)\left[1-P(0)+\int_{0}^{t} d s f(s) e^{-r s}\right]}_{\text {protection leg }}
$$

$$
\underbrace{(1-R)\left[1-P(0)+\int_{0}^{t} d s f(s) e^{-r s}\right]}_{\text {protection leg }} \stackrel{!}{=} \underbrace{c \int_{0}^{t} d s P(s) e^{-r s}}_{\text {premium leg }}
$$

$$
\underbrace{(1-R)\left[1-P(0)+\int_{0}^{t} d s f(s) e^{-r s}\right]}_{\text {protection leg }} \stackrel{!}{=} \underbrace{c \int_{0}^{t} d s P(s) e^{-r s}}_{\text {premium leg }}
$$

$$
\Rightarrow c=r(1-R) \frac{1-P(0)+e^{r \xi}(G(t+\xi)-G(\xi))}{P(0)-P(t) e^{-r t}-e^{r \xi}(G(t+\xi)-G(\xi))}
$$

where $\xi=\frac{\lambda^{2}}{\sigma^{2}}$ and $R$ is the expected recovery rate to a specific debt class

## Asset volatility estimation

## B. Appendix

## Asset volatility estimation

$\triangleright$ In the Merton model, it follows from Ito's lemma that

$$
\sigma_{S}=\frac{V}{S} \underbrace{\frac{\partial S}{\partial V}}_{0 \leq \ldots \leq 1} \sigma_{V} \Leftrightarrow \sigma_{V}=\frac{S}{V} \underbrace{\frac{\partial V}{\partial S}}_{\geq 1} \sigma_{S}
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$$

$\triangleright$ CreditGrades:

$$
\begin{array}{ll}
V=S+\bar{L} D & \text { and the following approximation } \\
\frac{\partial S}{\partial V}=1 & \text { result from boundary conditions } \\
\sigma_{S} \approx \sigma_{S}^{i m p} & \text { ATM implied volatility }
\end{array}
$$

$$
\Rightarrow \sigma_{S}=\frac{S+\bar{L} D}{S} \sigma_{V} \quad \Leftrightarrow \quad \sigma_{v}=\frac{S}{S+\bar{L} D} \sigma_{s}
$$

$\triangleright$ In our boundary examinations we focus on the distance to default since its behavior is relevant for determining the survival probability:

$$
\eta=\frac{1}{\sigma} \log \left(\frac{V}{L D}\right)=\frac{V}{\sigma_{S} S} \frac{\partial S}{\partial V} \log \left(\frac{V}{L D}\right)
$$

$\triangleright$ First (at/near to default) boundary condition

- Assume that as default approaches, $S \rightarrow 0$
- Thus at the boundary, $\left.V\right|_{s=0}=L D$, and near the boundary $V \approx L D+\frac{\partial V}{\partial S} S$
- Substituting into $\eta$ gives $\eta \approx \frac{1}{\sigma_{S}}\left(1+\frac{\partial V}{\partial S} \frac{S}{L D}\right) \approx \frac{1}{\sigma_{S}}$
$\triangleright$ Second (far from default) boundary condition
- Assume that as $V$ goes to infinity, $\frac{S}{V} \rightarrow 1$, i.e. $V$ and $S$ increase at the same rate, $\frac{\partial S}{\partial V} \rightarrow 1$
- Substitution leads to $\eta \approx \frac{1}{\sigma_{S}} \log \left(\frac{S}{L D}\right)$
$\triangleright$ The simplest expressions satisfying both boundary conditions are:

$$
\eta=\frac{S+L D}{\sigma_{S} S} \log \left(\frac{S+L D}{L D}\right) \quad \text { and } \quad V=S+L D
$$

$$
\sigma_{S}=\frac{V}{S} \frac{\partial S}{\partial V} \sigma_{V} \Rightarrow \sigma_{S}=\frac{S+\bar{L} D}{S} \sigma_{V}
$$

How sensitive are our results to $\frac{\partial S}{\partial V}<1$ ?
$\Rightarrow \sigma_{V} \nearrow$ as $\frac{\partial V}{\partial S} \nearrow$, i.e. stock-implied credit spreads would be even higher!
Alternatives:
$\triangleright$ Obtain $\sigma_{V}$ from $P\left(S, t, B, \sigma_{V}\right)$, the price of an equity put option as a function of $\sigma_{V}$, which can be equated to the market price of a put (Finger and Stamicar (2006)). Our own test runs confirm their conclusion that the differences to the baseline approach are marginal.

- Iterative approach suggested by Crosbie and Bohn (2003) and Vassalou and Xing (2004) applies to strict Merton setup (in which default may not occur at any point in time):
$\diamond$ Using either the historical or implied stock volatility as initial value for $\sigma_{V}$ and applying the BS formula, one can infer a time series of asset values to calculate the historical asset volatility, which is used as input for the next iteration. The described procedure is repeated until the historical volatility estimates from consecutive iterations converge.
$\diamond$ Iterative approach was shown to provide hardly any improvement over the direct approach (Bharath and Shumway (2008)).
$\diamond$ Through our calibration over $\bar{L}$, we determine $V$ and $\sigma_{V}$ simultaneously to be consistent with market observations.
$\triangleright$ Implied equity volatility is approximately an average of local volatilities (Derman, Iraj, and Zou (1995)):

$$
\sigma_{S}^{i m p} \approx \frac{1}{X-S} \int_{S}^{x} \sigma_{S} d S
$$

$\triangleright$ Substituting the local relation $\sigma_{S}=\sigma\left(1+\frac{B}{S}\right)$

$$
\sigma_{S}^{i m p} \approx \sigma\left\{1+\frac{B}{X-S} \log \left(\frac{X}{S}\right)\right\}
$$

$\triangleright$ At the money, i.e. for $S \rightarrow X$,

$$
\sigma_{s}^{i m p} \approx \sigma\left\{1+B \lim _{s \rightarrow X} \frac{\log \left(\frac{x}{s}\right)}{X-S}\right\}
$$

$\triangleright$ Applying l'Hôpital's rule gives

$$
\sigma_{S}^{i m p} \approx \sigma \frac{S+B}{S} \approx \sigma_{S}
$$

$\triangleright$ The standard deviation of the adjustment factor $L, \lambda$, is set to 0.3 (Finger et al. (2002)).
$\triangleright$ The debt class specific recovery rate, $R$, is set to 0.5 (Yu (2006)).
$\triangleright$ The debt per share, $D$, is calculated as $\frac{\text { total liabilities }}{\# \text { common shares outstanding }}$.
$\triangleright$ The risk-free interest rate, $r$, is assumed to be the five-year constant maturity zero-coupon swap rate inferred from swap rates.
$\triangleright$ The equity volatility, $\sigma_{S}$, is the one-year at-the-money implied volatility from put options.
$\triangleright$ Apply the Act/360 day counting convention.

## B. Appendix

## Constant default barrier

$\triangleright$ Determine $\bar{L}_{i}$ by minimizing the sum of squared errors between model ( $\overline{C D S}$ ) and market spreads (CDS) over a number of observations $N$ in the period 01/2003-07/2007:

$$
\min _{\bar{L}_{i}} \sum_{n=1}^{N}\left(\widehat{C D S}_{i, n}\left(\bar{L}_{i}\right)-C D S_{i, n}\right)^{2}
$$

## B. Appendix

$\triangleright$ Determine $\bar{L}_{i}$ by minimizing the sum of squared errors between model ( $\overline{C D S}$ ) and market spreads (CDS) over a number of observations $N$ in the period 01/2003-07/2007:

$$
\min _{\bar{L}_{i}} \sum_{n=1}^{N}\left(\widehat{C D S}_{i, n}\left(\bar{L}_{i}\right)-C D S_{i, n}\right)^{2}
$$

| Ival | Whole Sample Period |  |  |  |  | Pre-Crisis Period |  | Crisis Period |  | Post-Crisis Period |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | $\bar{N}$ | $\overline{\bar{L}}$ | cRMSE | ME | RMSE | ME | RMSE |  |  | ME | RMSE |
| 50 | 16 | 1.053 | 40.97 | 20.14 | 159.70 | -9.17 | 46.92 | 68.38 | 246.92 | 30.48 | 141.67 |
| 10 | 76 | 1.070 | 39.80 | 20.60 | 158.14 | -8.90 | 44.79 | 69.16 | 246.30 | 31.05 | 138.96 |
| 3 | 253 | 1.076 | 39.35 | 20.47 | 158.40 | -8.84 | 44.71 | 68.85 | 246.89 | 30.54 | 138.76 |
| 1 | 757 | 1.077 | 38.93 | 19.94 | 158.77 | -8.93 | 44.72 | 67.84 | 247.60 | 29.27 | 138.75 |

## B. Appendix

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| Ival | Whole Sample Period |  |  |  |  | Pre-Crisis Period |  | Crisis Period |  | Post-Crisis Period |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | $\bar{N}$ | $\bar{L}$ | cRMSE | ME | RMSE | ME | RMSE | ME | RMSE | ME | RMSE |
| 50 | 16 | 1.053 | 40.97 | 20.14 | 159.70 | -9.17 | 46.92 | 68.38 | 246.92 | 30.48 | 141.67 |
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$\triangleright$ Results very robust to choice of grid density. Reducing the interval from 50 to 10 slightly improves the estimates, therefore, focus on an interval of 10 in the following.
$\triangleright$ Determine $\bar{L}_{i}$ by minimizing the sum of squared errors between model $(\widehat{C D S})$ and market spreads (CDS) over a number of observations $N$ in the period 01/2003-07/2007:

$$
\min _{\bar{L}_{i}} \sum_{n=1}^{N}\left(\widehat{C D S}_{i, n}\left(\bar{L}_{i}\right)-C D S_{i, n}\right)^{2}
$$

| Ival | Whole Sample Period |  |  |  |  | Pre-Crisis Period |  | Crisis Period ME RMSE |  | Post-Crisis Period ME RMSE |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | $\bar{N}$ | $\bar{L}$ | cRMSE | ME | RMSE | ME | RMSE |  |  |  |  |
| 50 | 16 | 1.053 | 40.97 | 20.14 | 159.70 | -9.17 | 46.92 | 68.38 | 246.92 | 30.48 | 141.67 |
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| 1 | 757 | 1.077 | 38.93 | 19.94 | 158.77 | -8.93 | 44.72 | 67.84 | 247.60 | 29.27 | 138.75 |

$\triangleright$ Results very robust to choice of grid density. Reducing the interval from 50 to 10 slightly improves the estimates, therefore, focus on an interval of 10 in the following.
$\triangleright$ In the pre-crisis period the model underpredicts observed spreads due to nondefault components, like illiquidity, in line with the literature (Eom, Helwege, and Huang (2004), Longstaff (2004), Tang and Yan (2007)).

## B. Appendix

## Time-varying default barrier

$\triangleright$ Determine $\bar{L}_{i, t}$ daily by minimizing the sum of squared errors between model $(\widehat{C D S})$ and market spreads (CDS) based on a trailing window (with $N=5$ and an interval between calibration points $=2$ ):

$$
\min _{\bar{L}_{i, t}} \sum_{n=1}^{N}\left(\widehat{C D S}_{i, n}\left(\bar{L}_{i, t}\right)-C D S_{i, n}\right)^{2}
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\min _{\bar{L}_{i, t}} \sum_{n=1}^{N}\left(\widehat{C D S}_{i, n}\left(\bar{L}_{i, t}\right)-C D S_{i, n}\right)^{2}
$$

|  | Whole Sample Period |  |  |  |  | Pre-Crisis Period |  |  | Crisis Period |  |  | Post-Crisis Period |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | $\beta$ | p-value | $\bar{L}$ | ME | RMSE | $\bar{L}$ | ME | RMSE | $\bar{L}$ | ME | RMSE | $\overline{\bar{L}}$ | ME | RMSE |
| All | $-7.95 \mathrm{E}+07$ | 0.97 | 1.133 | -4.10 | 48.84 | 1.284 | -0.87 | 11.35 | 0.935 | -9.11 | 74.07 | 1.081 | -3.36 | 37.48 |
| Fin | -0.0004991 | 0.01 | 0.549 | -1.92 | 76.67 | 0.616 | -1.75 | 7.46 | 0.455 | 2.80 | 111.89 | 0.524 | -14.02 | 129.38 |
| Nonfin | $-1.57 \mathrm{E}+08$ | 0.95 | 1.232 | -4.47 | 43.98 | 1.402 | -0.71 | 12.03 | 1.013 | -11.05 | 67.47 | 1.168 | -1.68 | 22.08 |

$\triangleright$ The default boundary generally lowers during the crisis and slopes upwards in economic recovery without necessarily closing up to pre-crisis levels.

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## Time-varying default barrier

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$$

|  | Whole Sample Period |  |  |  |  | Pre-Crisis Period |  |  | Crisis Period |  |  | Post-Crisis Period |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | $\beta$ | p-value | $\overline{\bar{L}}$ | ME | RMSE | $\bar{L}$ | ME | RMSE | $\overline{\bar{L}}$ | ME | RMSE | $\bar{\tau}$ | ME | RMSE |
| All | -7.95E+07 | 0.97 | 1.133 | -4.10 | 48.84 | 1.284 | -0.87 | 11.35 | 0.935 | -9.11 | 74.07 | 1.081 | -3.36 | 37.48 |
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$\triangleright$ The average percentage decrease of $\bar{L}$ is about $25 \%$ over all sectors.

## B. Appendix

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$$
\min _{L_{i, t}} \sum_{n=1}^{N}\left(\widehat{C D S}_{i, n}\left(\bar{L}_{i, t}\right)-C D S_{i, n}\right)^{2}
$$

|  | Whole Sample Period |  |  |  |  | Pre-Crisis Period |  |  | Crisis Period |  |  | Post-Crisis Period |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | $\beta$ | p-value | $\overline{\bar{L}}$ | ME | RMSE | $\bar{L}$ | ME | RMSE | $\overline{\bar{L}}$ | ME | RMSE | $\bar{\tau}$ | ME | RMSE |
| All | -7.95E+07 | 0.97 | 1.133 | -4.10 | 48.84 | 1.284 | -0.87 | 11.35 | 0.935 | -9.11 | 74.07 | 1.081 | -3.36 | 37.48 |
| Fin | -0.0004991 | 0.01 | 0.549 | -1.92 | 76.67 | 0.616 | -1.75 | 7.46 | 0.455 | 2.80 | 111.89 | 0.524 | -14.02 | 129.38 |
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$\triangleright$ The default boundary generally lowers during the crisis and slopes upwards in economic recovery without necessarily closing up to pre-crisis levels.
$\triangleright$ The average percentage decrease of $\bar{L}$ is about $25 \%$ over all sectors.
$\triangleright$ However, a trend regression of daily percentage changes of $\bar{L}$ against time points $t$ reveals a significant negative trend only in the case of financials, not for the other companies.

## B. Appendix

## Time-varying default barrier



## B. Appendix

## Default barrier with a regime shift

$\triangleright$ The level of $\bar{L}$ can change exactly once from $\bar{L}_{1}$ to $\bar{L}_{2}$ at split date $t_{2}$. The estimation window ranges from 01/2004-12/2009 with a grid interval of 10. The minimization problem under these assumptions becomes:

$$
\left.\left.\bar{L}_{i, 1} \bar{m}_{i, 2}, t_{i, 2} \sum_{n=1}^{N} \widehat{(C D S}_{i, n}\left(\bar{L}_{i, 1}\right)-C D S_{i, n}\right)^{2} I_{\left\{\tau_{i, n}<t_{i, 2}\right\}}+\widehat{(C D S}_{i, n}\left(\bar{L}_{i, 2}\right)-C D S_{i, n}\right)^{2} I_{\left\{\tau_{i, n} \geq t_{i, 2}\right\}}
$$

|  | Whole Sample Period |  |  |  |  | Pre-Crisis Period ME RMSE |  | Crisis Period ME RMSE |  | Post-Crisis Period ME RMSE |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | $\overline{\bar{L}}_{1}$ | $\overline{\bar{L}}_{2}$ | Median $t_{2}$ | ME | RMSE |  |  |  |  |  |  |
| All | 1.056 | 0.920 | 09/30/2008 | -14.84 | 91.96 | -6.68 | 53.74 | -10.01 | 110.24 | -60.63 | 125.71 |
| Fin | 0.465 | 0.246 | 11/04/2008 | -26.16 | 124.73 | -16.86 | 39.84 | -21.41 | 171.21 | -81.82 | 258.29 |
| Nonfin | 1.159 | 1.038 | 09/30/2008 | -12.90 | 86.24 | -4.84 | 56.17 | -8.14 | 99.60 | -57.30 | 103.43 |

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$$
\left.\left.\bar{L}_{i, 1} \underline{L}_{i, 2}, t_{i, 2} \sum_{n=1}^{N} \widehat{(C D S}_{i, n}\left(\bar{L}_{i, 1}\right)-C D S_{i, n}\right)^{2} I_{\left\{\tau_{i, n}<t_{i, 2}\right\}}+\widehat{(C D S}_{i, n}\left(\bar{L}_{i, 2}\right)-C D S_{i, n}\right)^{2} I_{\left\{\tau_{i, n} \geq t_{i, 2}\right\}}
$$

|  | Whole Sample Period |  |  |  |  | Pre-Crisis Period ME RMSE |  | Crisis Period <br> ME RMSE |  | Post-Crisis Period |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | $\overline{\bar{L}}_{1}$ | $\overline{\bar{L}}_{2}$ | Median $t_{2}$ | ME | RMSE |  |  | ME | RMSE |
| All | 1.056 | 0.920 | 09/30/2008 | -14.84 | 91.96 | -6.68 | 53.74 |  |  | -10.01 | 110.24 | -60.63 | 125.71 |
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$\triangleright$ The median split date falls well within the tumultuous period following the bankruptcy of Lehman Brothers on 09/15/2008.
$\triangleright$ The level of $\bar{L}$ can change exactly once from $\bar{L}_{1}$ to $\bar{L}_{2}$ at split date $t_{2}$. The estimation window ranges from 01/2004-12/2009 with a grid interval of 10. The minimization problem under these assumptions becomes:

$$
\left.\left.\overline{m i n}_{i, 1}, \bar{L}_{i, 2}, t_{i, 2} \sum_{n=1}^{N} \widehat{(C D S}_{i, n}\left(\bar{L}_{i, 1}\right)-\operatorname{CDS}_{i, n}\right)^{2} I_{\left\{\tau_{i, n}<t_{i, 2}\right\}}+\widehat{(C D S}_{i, n}\left(\bar{L}_{i, 2}\right)-C D S_{i, n}\right)^{2} I_{\left\{\tau_{i, n} \geq t_{i, 2}\right\}}
$$


$\triangleright$ The median split date falls well within the tumultuous period following the bankruptcy of Lehman Brothers on 09/15/2008.
$\triangleright$ The post-crisis period is poorly fitted with negative mean errors, suggesting a second upward regime shift around mid 2009.

## B. Appendix

## Default barrier with a regime shift



## B. Appendix

Term structure of deviations


## B. Appendix

Term structure of deviations

$\triangleright$ The anticipation of bailouts matters most to short-term investors.

